

This supplemental document first describes the derivation of the variance $V[\langle F \rangle_{\text{tsr}}]$ for the (one-sample) two-stage resampling estimator in Eq. (12) of the main paper. Then the derivation of the variance $V[\langle I_t \rangle_{\text{tsr}}^N]$ for the (N -sample) two-stage resampling estimator for resampling strategies t is described. Finally, we derive the weighting function that aims to reduce the variance $V[\langle I_t \rangle_{\text{tsr}}^N]$.

1. Derivation of variance $V[\langle F \rangle_{\text{tsr}}]$

To derive the variance of (one-sample) two-stage resampling estimator $V[\langle F \rangle_{\text{tsr}}] = E[\langle F^2 \rangle_{\text{tsr}}] - E[\langle F \rangle_{\text{tsr}}]^2$, we first derive the expected value $E[\langle F \rangle_{\text{tsr}}]$ and then the second moment $E[\langle F^2 \rangle_{\text{tsr}}]$ is described. Let \mathbf{X} be the set of M_1 proposals $\mathbf{X} = \{X_1, \dots, X_{M_1}\}$ sampled by the sampling pdf p , and \bar{X} be the subset of \mathbf{X} comprising of M_2 proposals $\bar{X} = \{X_{i_1}, \dots, X_{i_{M_2}}\}$. The expected value $E[\langle F \rangle_{\text{tsr}}]$ is calculated by:

$$E[\langle F \rangle_{\text{tsr}}] = E \left[\frac{f(\mathbf{X})}{\hat{q}_2(\mathbf{X})} \left(\frac{1}{M_1} \sum_{j=1}^{M_1} \frac{\hat{q}_1(X_j)}{p(X_j)} \right) \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \hat{q}_2(X_{i_j}) \right) \right]$$

By using $E[abc] = E[E[a|b]b|c]$, the expected value $E[\langle F \rangle_{\text{tsr}}]$ is represented as:

$$E \left[E \left[E \left[\frac{f(\mathbf{X})}{\hat{q}_2(\mathbf{X})} | \bar{X} \right] \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \hat{q}_2(X_{i_j}) \right) | \mathbf{X} \right] \left(\frac{1}{M_1} \sum_{j=1}^{M_1} \frac{\hat{q}_1(X_j)}{p(X_j)} \right) \right].$$

The innermost expected value is calculated by summing the product of f/\hat{q}_2 and the pmf over all the proposals in \bar{X} :

$$\begin{aligned} E \left[\frac{f(\mathbf{X})}{\hat{q}_2(\mathbf{X})} | \bar{X} \right] &= \sum_{j=1}^{M_2} \frac{f(X_{i_j})}{\hat{q}_2(X_{i_j})} \frac{\hat{q}_2(X_{i_j})/\hat{q}_1(X_{i_j})}{\sum_{k=1}^{M_2} \hat{q}_2(X_{i_k})/\hat{q}_1(X_{i_k})} \\ &= \left(\sum_{j=1}^{M_2} \frac{f(X_{i_j})}{\hat{q}_1(X_{i_j})} \right) \left(\sum_{k=1}^{M_2} \frac{\hat{q}_2(X_{i_k})}{\hat{q}_1(X_{i_k})} \right)^{-1}. \end{aligned}$$

By substituting this, $E[\langle F \rangle_{\text{tsr}}]$ is expressed as:

$$\begin{aligned} E \left[E \left[\frac{1}{M_2} \left(\sum_{j=1}^{M_2} \frac{f(X_{i_j})}{\hat{q}_1(X_{i_j})} \right) | \mathbf{X} \right] \left(\frac{1}{M_1} \sum_{j=1}^{M_1} \frac{\hat{q}_1(X_j)}{p(X_j)} \right) \right], \\ = E \left[E \left[\frac{f(X_{i_1})}{\hat{q}_1(X_{i_1})} | \mathbf{X} \right] \left(\frac{1}{M_1} \sum_{j=1}^{M_1} \frac{\hat{q}_1(X_j)}{p(X_j)} \right) \right]. \end{aligned}$$

The inner expected value is calculated by summing the product of f/\hat{q}_1 and the pmf over all the proposals in \mathbf{X} :

$$\begin{aligned} E \left[\frac{f(X_{i_1})}{\hat{q}_1(X_{i_1})} | \mathbf{X} \right] &= \sum_{i=1}^{M_1} \frac{f(X_i)}{\hat{q}_1(X_i)} \frac{\hat{q}_1(X_i)/p(X_i)}{\sum_{k=1}^{M_1} \hat{q}_1(X_k)/p(X_k)} \\ &= \left(\sum_{i=1}^{M_1} \frac{f(X_i)}{p(X_i)} \right) \left(\sum_{k=1}^{M_1} \frac{\hat{q}_1(X_k)}{p(X_k)} \right)^{-1}. \end{aligned}$$

By substituting this, the expected value $E[\langle F \rangle_{\text{tsr}}]$ is simplified to:

$$E \left[\frac{1}{M_1} \sum_{i=1}^{M_1} \frac{f(X_i)}{p(X_i)} \right] = E \left[\frac{f(\mathbf{X})}{p(\mathbf{X})} \right]. \quad (1)$$

Next, the second moment $E[\langle F^2 \rangle_{\text{tsr}}]$ is calculated as:

$$E \left[E \left[E \left[\left(\frac{f(\mathbf{X})}{\hat{q}_2(\mathbf{X})} \right)^2 | \bar{X} \right] \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{\hat{q}_2(X_{i_j})}{\hat{q}_1(X_{i_j})} \right)^2 | \mathbf{X} \right] \left(\frac{1}{M_1} \sum_{j=1}^{M_1} \frac{\hat{q}_1(X_j)}{p(X_j)} \right)^2 \right].$$

We first expand the innermost expected value as:

$$\begin{aligned} E \left[\left(\frac{f(\mathbf{X})}{\hat{q}_2(\mathbf{X})} \right)^2 | \bar{X} \right] &= \sum_{j=1}^{M_2} \frac{f^2(X_{i_j})}{\hat{q}_2^2(X_{i_j})} \frac{\hat{q}_2(X_{i_j})/\hat{q}_1(X_{i_j})}{\sum_{k=1}^{M_2} \hat{q}_2(X_{i_k})/\hat{q}_1(X_{i_k})} \\ &= \left(\sum_{j=1}^{M_2} \frac{f^2(X_{i_j})}{\hat{q}_1(X_{i_j})\hat{q}_2(X_{i_j})} \right) \left(\sum_{k=1}^{M_2} \frac{\hat{q}_2(X_{i_k})}{\hat{q}_1(X_{i_k})} \right)^{-1}. \end{aligned}$$

By substituting this, the inner expected value is transformed as:

$$\begin{aligned} E \left[\left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{f^2(X_{i_j})}{\hat{q}_1(X_{i_j})\hat{q}_2(X_{i_j})} \right) \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \hat{q}_2(X_{i_j}) \right) \middle| \mathbf{X} \right] \\ = \frac{1}{M_2} E \left[\frac{f^2}{\hat{q}_1^2} \middle| \mathbf{X} \right] + \left(1 - \frac{1}{M_2} \right) E \left[\frac{f^2}{\hat{q}_1\hat{q}_2} \middle| \mathbf{X} \right] E \left[\frac{\hat{q}_2}{\hat{q}_1} \middle| \mathbf{X} \right], \end{aligned}$$

These expected values are expanded as:

$$\begin{aligned} E \left[\frac{f^2(X_{i_1})}{\hat{q}_1^2(X_{i_1})} \middle| \mathbf{X} \right] &= \left(\sum_{j=1}^{M_1} \frac{f^2(X_j)}{\hat{q}_1(X_j)p(X_j)} \right) \left(\sum_{k=1}^{M_1} \frac{\hat{q}_1(X_k)}{p(X_k)} \right)^{-1} \\ E \left[\frac{f^2(X_{i_1})}{\hat{q}_1(X_{i_1})\hat{q}_2(X_{i_1})} \middle| \mathbf{X} \right] &= \left(\sum_{j=1}^{M_1} \frac{f^2(X_j)}{\hat{q}_2(X_j)p(X_j)} \right) \left(\sum_{k=1}^{M_1} \frac{\hat{q}_1(X_k)}{p(X_k)} \right)^{-1} \\ E \left[\frac{\hat{q}_2(X_{i_1})}{\hat{q}_1(X_{i_1})} \middle| \mathbf{X} \right] &= \left(\sum_{j=1}^{M_1} \frac{\hat{q}_2(X_j)}{p(X_j)} \right) \left(\sum_{k=1}^{M_1} \frac{\hat{q}_1(X_k)}{p(X_k)} \right)^{-1} \end{aligned}$$

By substituting these, the outermost expected value is expressed as:

$$\begin{aligned} \frac{1}{M_2} E \left[\left(\frac{1}{M_1} \sum_{j=1}^{M_1} \frac{f^2(X_j)}{\hat{q}_1(X_j)p(X_j)} \right) \left(\frac{1}{M_1} \sum_{k=1}^{M_1} \frac{\hat{q}_1(X_k)}{p(X_k)} \right) \right] \\ + \left(1 - \frac{1}{M_2} \right) E \left[\left(\frac{1}{M_1} \sum_{j=1}^{M_1} \frac{f^2(X_j)}{\hat{q}_2(X_j)p(X_j)} \right) \left(\frac{1}{M_1} \sum_{k=1}^{M_1} \frac{\hat{q}_2(X_k)}{p(X_k)} \right) \right] \\ = \frac{1}{M_1} E \left[\frac{f^2(X)}{p^2(X)} \right] + \left(1 - \frac{1}{M_1} \right) \frac{1}{M_2} E \left[\frac{f^2(X)}{\hat{q}_1(X)p(X)} \right] E \left[\frac{\hat{q}_1(X)}{p(X)} \right] \\ + \left(1 - \frac{1}{M_1} \right) \left(1 - \frac{1}{M_2} \right) E \left[\frac{f^2(X)}{\hat{q}_2(X)p(X)} \right] E \left[\frac{\hat{q}_2(X)}{p(X)} \right]. \end{aligned}$$

We omit the arguments from here on for the brevity. Then the product of two expected values $E[f^2/\hat{q}_1p]E[\hat{q}_1/p]$ is rewritten as:

$$\begin{aligned} E \left[\frac{f^2}{\hat{q}_1p} \right] E \left[\frac{\hat{q}_1}{p} \right] &= \int \frac{f^2(x)}{\hat{q}_1(x)} dx \cdot \int \hat{q}_1(x) dx \\ &= \int \frac{f^2(x)}{\hat{q}_1(x)/\int \hat{q}_1(x) dx} dx = \int \frac{f^2(x)}{q_1(x)} dx = E \left[\frac{f^2}{q_1} \right], \end{aligned}$$

where q_1 is the target pdf of the target distribution \hat{q}_1 . Analogously, the product of two expected values $E[f^2/\hat{q}_2p]E[\hat{q}_2/p]$ is rewritten by using the target pdf q_2 as:

$$E \left[\frac{f^2}{\hat{q}_2p} \right] E \left[\frac{\hat{q}_2}{p} \right] = E \left[\frac{f^2}{q_2} \right].$$

By substituting these, $E[\langle F^2 \rangle_{\text{tsr}}]$ is expressed as:

$$\frac{1}{M_1} E \left[\frac{f^2}{p^2} \right] + \left(1 - \frac{1}{M_1} \right) \left(\frac{1}{M_2} E \left[\frac{f^2}{q_1} \right] + \left(1 - \frac{1}{M_2} \right) E \left[\frac{f^2}{q_2} \right] \right).$$

Finally, the variance $V[\langle F \rangle_{\text{tsr}}] = E[\langle F^2 \rangle_{\text{tsr}}] - E[\langle F \rangle_{\text{tsr}}]^2$ is calculated by subtracting the square of the expected value from the second moment. Here, the expected value can be rewritten by using the target pdfs q_1 and q_2 as $E[f/p] = \int f(x) dx = E[f/q_1] = E[f/q_2]$, and the sum of the coefficients in the second moment is equal to one (i.e., $1/M_1 + (1 - M_1)(1/M_2 + (1 - 1/M_2)) = 1$). Then the square of the expected value $E[f/p]^2$ can be expressed as:

$$\frac{1}{M_1} E \left[\frac{f}{p} \right]^2 + \left(1 - \frac{1}{M_1} \right) \left(\frac{1}{M_2} E \left[\frac{f}{q_1} \right]^2 + \left(1 - \frac{1}{M_2} \right) E \left[\frac{f}{q_2} \right]^2 \right).$$

By subtracting these from the second moment, the variance is represented by Eq. (12) of the main paper.

2. Derivation of Variance for Two-Stage Resampling Estimator

In this supplemental material, we explain the derivation of the variance for (N-sample) two-stage resampling estimator $V[\langle I_t \rangle_{\text{tsr}}^N]$. As described in the paper, we assume that the importance sampling is used (instead of the stratified sampling) in the first resampling stage during the derivation of the variance. This is a reasonable assumption since the variance of the stratified sampling is smaller than that of the importance sampling in general. By using this assumption, the two-stage resampling estimator $V[\langle I_t \rangle_{\text{tsr}}^N]$ for a given eye sub-path sample \bar{Z}_t is calculated from:

$$\langle I_t \rangle_{\text{tsr}}^N = \frac{1}{M_1 M_2 N} \sum_{k=1}^N \frac{w_t(\bar{Y}_k \bar{Z}_t) f(\bar{Y}_k \bar{Z}_t)}{p(\bar{Y}_k) p(\bar{Z}_t) P_r(\bar{Y}_k | \bar{\mathbf{Y}})},$$

where M_1 is the number of tracing light sub-paths, \bar{Y} is the set of pre-sampled light sub-paths as $\bar{\mathbf{Y}} = \{\bar{Y}_{1,1}, \dots, \bar{Y}_{s,i}, \dots\}$ where $\bar{Y}_{s,i}$ is i -th light sub-path sample with s vertices. M_2 is the number of light sub-paths for the second resampling stage, and the subset of pre-sampled light sub-paths used in the second resampling stage is represented by $\bar{\mathbf{Y}} = \{\bar{Y}_{i_1}, \dots, \bar{Y}_{i_j}, \dots, \bar{Y}_{i_{M_2}}\}$ where $\bar{Y}_{i_j} \in \bar{\mathbf{Y}}$. \bar{Y}_k is the k -th sample resampled from the subset. To simplify the notation, we concatenate the light sub-path with the sampled eye sub-path \bar{Z}_t , and the light sub-paths such as \bar{Y}_k , $\bar{Y}_{s,i}$, and \bar{Y}_{i_j} are represented by the full light path samples as $\bar{X}_k = \bar{Y}_k \bar{Z}_t$, $\bar{X}_{s,i} = \bar{Y}_{s,i} \bar{Z}_t$, and $\bar{X}_{i_j} = \bar{Y}_{i_j} \bar{Z}_t$. We also represent the measurement contribution function f weighted by the weighting function w_t with $f_t(\bar{x}) \equiv w_t(\bar{x}) f(\bar{x})$, and the set $\bar{\mathbf{Y}}$ is replaced with $\bar{\mathbf{X}} = \{\bar{X}_{1,1}, \dots, \bar{X}_{s,i}, \dots\}$. The subset used in the second resampling stage is denoted by $\bar{\mathbf{X}} = \{\bar{X}_{i_1}, \dots, \bar{X}_{i_{M_2}}\}$. By using these notations, the two-stage resampling estimator $\langle I_t \rangle_{\text{tsr}}^N$ is simplified to:

$$\langle I_t \rangle_{\text{tsr}}^N = \frac{1}{M_1 M_2 N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{p(\bar{X}_k) P_r(\bar{X}_k | \bar{\mathbf{X}})}, \quad (2)$$

The resampling pmf P_r is the product of two pmfs as:

$$P_1(\bar{X}_k | \bar{\mathbf{X}}) = \frac{\hat{q}_1(\bar{X}_k) / p(\bar{X}_k)}{\sum_{s \geq 1} \sum_{i=1}^{M_1} \hat{q}_1(\bar{X}_{s,i}) / p(\bar{X}_{s,i})}, \quad (3)$$

$$P_2(\bar{X}_k | \bar{\mathbf{X}}) = \frac{\hat{q}_2(\bar{X}_k) / \hat{q}_1(\bar{X}_k)}{\sum_{j=1}^{M_2} \hat{q}_2(\bar{X}_{i_j}) / \hat{q}_1(\bar{X}_{i_j})}, \quad (4)$$

where \hat{q}_1 and \hat{q}_2 are the target distributions for the first resampling stage and the second resampling stage, respectively. By substituting P_1 and P_2 for P_r , $\langle I_t \rangle_{\text{tsr}}^N$ is expressed by:

$$\begin{aligned} \langle I_t \rangle_{\text{tsr}}^N &= \frac{1}{M_1 M_2 N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{p(\bar{X}_k) P_1(\bar{X}_k | \bar{\mathbf{X}}) P_2(\bar{X}_k | \bar{\mathbf{X}})} \\ &= \left(\frac{1}{N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{\hat{q}_2(\bar{X}_k)} \right) \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right) \end{aligned} \quad (5)$$

Similarly, $\langle I_t^2 \rangle_{\text{tsr}}^N$ used for the second moment $E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is represented by:

$$\langle I_t^2 \rangle_{\text{tsr}}^N = \left(\frac{1}{N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{\hat{q}_2(\bar{X}_k)} \right)^2 \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^2 \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right)^2. \quad (6)$$

To compute the conditional variance $V[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t] = E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t] - E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]^2$, we first explain the derivation of $E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$, then the derivation of $E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is described.

2.1. Derivation of $E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$

The conditional expected value $E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is expressed by:

$$E \left[\left(\frac{1}{N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{\hat{q}_2(\bar{X}_k)} \right) \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right) \middle| \bar{Z}_t \right]. \quad (7)$$

By using $E[abc] = E[E[a|b]b|c]$, $E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is transformed into:

$$E \left[E \left[E \left[\left(\frac{1}{N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{\hat{q}_2(\bar{X}_k)} \right) \middle| \bar{\mathbf{X}} \right] \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right) \middle| \bar{\mathbf{X}} \right] \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) \middle| \bar{Z}_t \right].$$

We first expand the innermost expected value as follows:

$$E \left[\left(\frac{1}{N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{\hat{q}_2(\bar{X}_k)} \right) | \bar{\mathbf{X}} \right] = E \left[\frac{f_t(\bar{X}_1)}{\hat{q}_2(\bar{X}_1)} | \bar{\mathbf{X}} \right] = \sum_{j=1}^{M_2} \frac{f_t(\bar{X}_{i_j})}{\hat{q}_2(\bar{X}_{i_j})} \frac{\hat{q}_2(\bar{X}_{i_j})/\hat{q}_1(\bar{X}_{i_j})}{\sum_{k=1}^{M_2} \hat{q}_2(\bar{X}_{i_k})/\hat{q}_1(\bar{X}_{i_k})} = \left(\sum_{j=1}^{M_2} \frac{f_t(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right) \left(\sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right)^{-1} \quad (8)$$

Substituting the above equation cancels out the summation and leads to:

$$E \left[E \left[\left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{f_t(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right) | \bar{\mathbf{X}} \right] \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) | \bar{Z}_t \right] = E \left[E \left[\frac{f_t(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) | \bar{Z}_t \right] \quad (9)$$

The inner expected value of the above equation is calculated by summing over all proposals in $\bar{\mathbf{X}}$ as:

$$E \left[\frac{f_t(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] = \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t(\bar{X}_{s,i})}{\hat{q}_1(\bar{X}_{s,i})} \cdot \left(\frac{\hat{q}_1(\bar{X}_{s,i})/p(\bar{X}_{s,i})}{\sum_{s \geq 1} \sum_{i=1}^{M_1} \hat{q}_1(\bar{X}_{s,i})/p(\bar{X}_{s,i})} \right) = \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^{-1}$$

By substituting the above equation and canceling out the summation, $E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is expressed by:

$$E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t] = E \left[\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} | \bar{Z}_t \right] = \sum_{s \geq 1} E \left[\frac{1}{M_1} \sum_{i=1}^{M_1} \frac{f_t(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} | \bar{Z}_t \right] = \sum_{s \geq 1} E \left[\frac{f_t(\bar{X}_{s,1})}{p(\bar{X}_{s,1})} | \bar{Z}_t \right]. \quad (10)$$

2.2. Derivation of the second moment $E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t]$

Similar to the derivation of $E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$, by using $E[abc] = E[E[a|b]b|c]$, the conditional expected value $E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is transformed into:

$$E \left[E \left[E \left[\left(\frac{1}{N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{\hat{q}_2(\bar{X}_k)} \right)^2 | \bar{\mathbf{X}} \right] \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right)^2 | \bar{\mathbf{X}} \right] \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^2 | \bar{Z}_t \right].$$

By using the following identity:

$$E \left[\left(\frac{1}{N} \sum_{i=1}^N f(X_i) \right) \left(\frac{1}{N} \sum_{i=1}^N g(X_i) \right) \right] = \frac{1}{N} E[f(X_1)g(X_1)] + \left(1 - \frac{1}{N} \right) E[f(X_1)]E[g(X_1)],$$

the innermost expected value is expressed by:

$$\begin{aligned} E \left[\left(\frac{1}{N} \sum_{k=1}^N \frac{f_t(\bar{X}_k)}{\hat{q}_2(\bar{X}_k)} \right)^2 | \bar{\mathbf{X}} \right] &= \frac{1}{N} E \left[\frac{f_t^2(\bar{X}_1)}{\hat{q}_2^2(\bar{X}_1)} | \bar{\mathbf{X}} \right] + \left(1 - \frac{1}{N} \right) E \left[\frac{f_t(\bar{X}_1)}{\hat{q}_2(\bar{X}_1)} | \bar{\mathbf{X}} \right]^2 \\ &= \frac{1}{N} \sum_{j=1}^{M_2} \frac{f_t^2(\bar{X}_{i_j})}{\hat{q}_2^2(\bar{X}_{i_j})} \frac{\hat{q}_2(\bar{X}_{i_j})/\hat{q}_1(\bar{X}_{i_j})}{\sum_{k=1}^{M_2} \hat{q}_2(\bar{X}_{i_k})/\hat{q}_1(\bar{X}_{i_k})} + \left(1 - \frac{1}{N} \right) \left(\sum_{j=1}^{M_2} \frac{f_t(\bar{X}_{i_j})}{\hat{q}_2(\bar{X}_{i_j})} \frac{\hat{q}_2(\bar{X}_{i_j})/\hat{q}_1(\bar{X}_{i_j})}{\sum_{k=1}^{M_2} \hat{q}_2(\bar{X}_{i_k})/\hat{q}_1(\bar{X}_{i_k})} \right)^2 \\ &= \frac{1}{N} \left(\sum_{j=1}^{M_2} \frac{f_t^2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})\hat{q}_2(\bar{X}_{i_j})} \right) \left(\sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right)^{-1} + \left(1 - \frac{1}{N} \right) \left(\sum_{j=1}^{M_2} \frac{f_t(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right)^2 \left(\sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right)^{-2}. \end{aligned}$$

By substituting this, the inner expected value leads to:

$$\begin{aligned} E \left[\frac{1}{N} \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{f_t^2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})\hat{q}_2(\bar{X}_{i_j})} \right) \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right) + \left(1 - \frac{1}{N} \right) \left(\frac{1}{M_2} \sum_{j=1}^{M_2} \frac{f_t(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \right)^2 | \bar{\mathbf{X}} \right] \\ = \frac{1}{N} \left(\frac{1}{M_2} E \left[\frac{f_t^2(\bar{X}_{i_1})}{\hat{q}_1^2(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] + \left(1 - \frac{1}{M_2} \right) E \left[\frac{f_t^2(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})\hat{q}_2(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] E \left[\frac{\hat{q}_2(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] \right) + \left(1 - \frac{1}{N} \right) \left(\frac{1}{M_2} E \left[\frac{f_t^2(\bar{X}_{i_1})}{\hat{q}_1^2(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] + \left(1 - \frac{1}{M_2} \right) E \left[\frac{f_t(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right]^2 \right) \\ = \frac{1}{M_2} E \left[\frac{f_t^2(\bar{X}_{i_1})}{\hat{q}_1^2(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] + \frac{1}{N} \left(1 - \frac{1}{M_2} \right) E \left[\frac{f_t^2(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})\hat{q}_2(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] E \left[\frac{\hat{q}_2(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right] + \left(1 - \frac{1}{N} \right) \left(1 - \frac{1}{M_2} \right) E \left[\frac{f_t(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})} | \bar{\mathbf{X}} \right]^2, \end{aligned}$$

where we use the following identity:

$$E \left[\left(\frac{1}{M} \sum_{s \geq 1} \sum_{i=1}^M f(X_{s,i}) \right) \left(\frac{1}{M} \sum_{s \geq 1} \sum_{i=1}^M g(X_{s,i}) \right) \right] = \frac{1}{M} \sum_{s \geq 1} E[f(X_{s,1})g(X_{s,1})] - \frac{1}{M} \sum_{s \geq 1} E[f(X_{s,1})]E[g(X_{s,1})] + \left(\sum_{s \geq 1} E[f(X_{s,1})] \right) \left(\sum_{s \geq 1} E[g(X_{s,1})] \right).$$

By expanding each expected value over all proposals in the set $\bar{\mathbf{X}}$, the above equation is expressed as:

$$\begin{aligned} & \frac{1}{M_2} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t^2(\bar{X}_{s,i})}{\hat{q}_1^2(\bar{X}_{s,i})} \frac{\hat{q}_1(\bar{X}_{s,i})/p(\bar{X}_{s,i})}{\sum_{s \geq 1} \sum_{j=1}^{M_1} \hat{q}_1(\bar{X}_{s,j})/p(\bar{X}_{s,j})} + \frac{1}{N} \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t^2(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})\hat{q}_2(\bar{X}_{i_1})} \frac{\hat{q}_1(\bar{X}_{s,i})/p(\bar{X}_{s,i})}{\sum_{s \geq 1} \sum_{j=1}^{M_1} \hat{q}_1(\bar{X}_{s,j})/p(\bar{X}_{s,j})} \right) \\ & \times \left(\sum_{s \geq 1} \sum_{j=1}^{M_1} \frac{\hat{q}_2(\bar{X}_{i_j})}{\hat{q}_1(\bar{X}_{i_j})} \frac{\hat{q}_1(\bar{X}_{s,i})/p(\bar{X}_{s,i})}{\sum_{s \geq 1} \sum_{j=1}^{M_1} \hat{q}_1(\bar{X}_{s,j})/p(\bar{X}_{s,j})} \right) + \left(1 - \frac{1}{N}\right) \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t(\bar{X}_{i_1})}{\hat{q}_1(\bar{X}_{i_1})} \frac{\hat{q}_1(\bar{X}_{s,i})/p(\bar{X}_{s,i})}{\sum_{s \geq 1} \sum_{j=1}^{M_1} \hat{q}_1(\bar{X}_{s,j})/p(\bar{X}_{s,j})} \right)^2 \\ & = \frac{1}{M_2} \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t^2(\bar{X}_{s,i})}{\hat{q}_1(\bar{X}_{s,i})p(\bar{X}_{s,i})} \right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^{-1} + \frac{1}{N} \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t^2(\bar{X}_{s,i})}{p(\bar{X}_{s,i})\hat{q}_2(\bar{X}_{s,i})} \right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_2(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^{-2} \\ & + \left(1 - \frac{1}{N}\right) \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^2 \left(\sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^{-2}. \end{aligned}$$

Substituting the above equation, $E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is expressed by:

$$\begin{aligned} & \frac{1}{M_2} E \left[\left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t^2(\bar{X}_{s,i})}{\hat{q}_1(\bar{X}_{s,i})p(\bar{X}_{s,i})} \right) \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_1(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) | \bar{Z}_t \right] + \frac{1}{N} \left(1 - \frac{1}{M_2}\right) E \left[\left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t^2(\bar{X}_{s,i})}{p(\bar{X}_{s,i})\hat{q}_2(\bar{X}_{s,i})} \right) \left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{\hat{q}_2(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right) | \bar{Z}_t \right] \\ & + \left(1 - \frac{1}{N}\right) \left(1 - \frac{1}{M_2}\right) E \left[\left(\frac{1}{M_1} \sum_{s \geq 1} \sum_{i=1}^{M_1} \frac{f_t(\bar{X}_{s,i})}{p(\bar{X}_{s,i})} \right)^2 | \bar{Z}_t \right]. \end{aligned}$$

Then $E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is expressed by (we omit the arguments for brevity):

$$\begin{aligned} & \frac{1}{M_2} \left(\frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t^2}{p^2} | \bar{Z}_t \right] - \frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_1} | \bar{Z}_t \right] E \left[\frac{\hat{q}_1}{p} | \bar{Z}_t \right] + \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_1} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_1}{p} | \bar{Z}_t \right] \right) \right) \\ & + \frac{1}{N} \left(1 - \frac{1}{M_2}\right) \left(\frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t^2}{p^2} | \bar{Z}_t \right] - \frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_2} | \bar{Z}_t \right] E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] + \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_2} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] \right) \right) \\ & + \left(1 - \frac{1}{N}\right) \left(1 - \frac{1}{M_2}\right) \left(\frac{1}{M_1} E \left[\frac{f_t^2}{p^2} | \bar{Z}_t \right] - \frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t^2}{p} | \bar{Z}_t \right]^2 + \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_2} | \bar{Z}_t \right] \right)^2 \right). \end{aligned}$$

By rearranging the above equation, the conditional expected value $E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is expressed by:

$$\begin{aligned} E[\langle I_t^2 \rangle_{\text{tsr}}^N | \bar{Z}_t] & = \frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t^2}{p^2} | \bar{Z}_t \right] - \frac{1}{M_1 M_2} \sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_1} | \bar{Z}_t \right] E \left[\frac{\hat{q}_1}{p} | \bar{Z}_t \right] + \frac{1}{M_2} \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_1} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_1}{p} | \bar{Z}_t \right] \right) \\ & - \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2}\right) \sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_2} | \bar{Z}_t \right] E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] + \frac{1}{N} \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p\hat{q}_2} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] \right) \\ & - \left(1 - \frac{1}{N}\right) \frac{1}{M_1} \left(1 - \frac{1}{M_2}\right) \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 + \left(1 - \frac{1}{N}\right) \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right] \right)^2. \end{aligned}$$

To transform the above equation into the form of the variance, we expand the coefficients of the last line as follows:

$$\begin{aligned} & - \left(1 - \frac{1}{N}\right) \frac{1}{M_1} \left(1 - \frac{1}{M_2}\right) \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 + \left(1 - \frac{1}{N}\right) \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right] \right)^2 \\ & = - \left(1 - \frac{1}{N} - \frac{1}{M_2} + \frac{1}{N} \frac{1}{M_2}\right) \frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 + \left(1 - \frac{1}{N} - \frac{1}{M_2} + \frac{1}{N} \frac{1}{M_2}\right) \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right] \right)^2 \\ & = - \frac{1}{M_1} \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 + \frac{1}{M_1 M_2} \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 + \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2}\right) \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 \\ & + \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right] \right)^2 - \frac{1}{M_2} \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right] \right)^2 - \frac{1}{N} \left(1 - \frac{1}{M_2}\right) \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right] \right)^2. \end{aligned}$$

2.3. Derivation of the conditional variance $V[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$

The conditional variance $V[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is calculated by subtracting $E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]^2 = (\sum_{s \geq 1} E[f_t/p | \bar{Z}_t])^2$ from the above equation. Rearranging each component leads to:

$$\begin{aligned} & \frac{1}{M_1} \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p^2} | \bar{Z}_t \right] - \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 \right) - \frac{1}{M_1 M_2} \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p \hat{q}_1} | \bar{Z}_t \right] E \left[\frac{\hat{q}_1}{p} | \bar{Z}_t \right] - \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 \right) \\ & + \frac{1}{M_2} \left(\left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p \hat{q}_1} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_1}{p} | \bar{Z}_t \right] \right) - \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 \right) \right) - \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p \hat{q}_2} | \bar{Z}_t \right] E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] - \sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 \right) \\ & + \frac{1}{N} \left(1 - \frac{1}{M_2} \right) \left(\left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p \hat{q}_2} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] \right) - \left(\sum_{s \geq 1} E \left[\frac{f_t}{p} | \bar{Z}_t \right]^2 \right) \right) \end{aligned}$$

We now transform the product of two expected values shown in the above equation as:

$$\begin{aligned} E \left[\frac{f_t^2(\bar{X}_{s,1})}{p(\bar{X}_{s,1}) \hat{q}_1(\bar{X}_{s,1})} | \bar{Z}_t \right] E \left[\frac{\hat{q}_1(\bar{X}_{s,1})}{p(\bar{X}_{s,1})} | \bar{Z}_t \right] &= \frac{1}{p^2(\bar{Z}_t)} \left(\int_{A^s} \frac{f_t^2(\bar{y}_s \bar{Z}_t)}{\hat{q}_1(\bar{y}_s \bar{Z}_t)} d\mu(\bar{y}_s) \right) \underbrace{\left(\int_{A^s} \hat{q}_1(\bar{y}_s \bar{Z}_t) d\mu(\bar{y}_s) \right)}_{Q_{1,s}} = \frac{1}{p^2(\bar{Z}_t)} \left(\int_{A^s} \frac{f_t^2(\bar{y}_s \bar{Z}_t)}{\hat{q}_1(\bar{y}_s \bar{Z}_t) / Q_{1,s}} d\mu(\bar{y}_s) \right) \\ &= \frac{1}{p^2(\bar{Z}_t)} \left(\int_{A^s} \frac{f_t^2(\bar{y}_s \bar{Z}_t)}{q_{1,s}(\bar{y}_s | \bar{Z}_t)} d\mu(\bar{y}_s) \right) = \frac{1}{p^2(\bar{Z}_t)} E \left[\frac{f_t^2(\bar{X}_{s,1})}{q_{1,s}^2(\bar{X}_{s,1} | \bar{Z}_t)} | \bar{Z}_t \right] = E \left[\frac{f_t^2(\bar{X}_{s,1})}{q_{1,s}^2(\bar{X}_{s,1})} | \bar{Z}_t \right], \end{aligned}$$

where $Q_{1,s}$ is the normalization factor that is the integral of the target distribution \hat{q}_1 over all the light sub-paths \bar{y}_s with s vertices, and $q_{1,s}(\bar{y}_s | \bar{Z}_t)$ is the conditional pdf, and $q_1(\bar{x}_s) = q_{1,s}(\bar{y}_s | \bar{Z}_t) p(\bar{Z}_t)$. Similarly, the following product of two expected values can be calculated as:

$$\begin{aligned} \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p \hat{q}_1} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_1}{p} | \bar{Z}_t \right] \right) &= \frac{1}{p^2(\bar{Z}_t)} \left(\int_{\mathcal{A}} \frac{f_t^2(\bar{y} \bar{Z}_t)}{\hat{q}_1(\bar{y} \bar{Z}_t)} d\mu(\bar{y}) \right) \underbrace{\left(\int_{\mathcal{A}} \hat{q}_1(\bar{y} \bar{Z}_t) d\mu(\bar{y}) \right)}_{Q_1} = \frac{1}{p^2(\bar{Z}_t)} \left(\int_{\mathcal{A}} \frac{f_t^2(\bar{y} \bar{Z}_t)}{\hat{q}_1(\bar{y} \bar{Z}_t) / Q_1} d\mu(\bar{y}) \right) \\ &= \frac{1}{p^2(\bar{Z}_t)} \left(\int_{\mathcal{A}} \frac{f_t^2(\bar{y} \bar{Z}_t)}{q_1(\bar{y} | \bar{Z}_t)} d\mu(\bar{y}) \right) = E \left[\frac{f_t^2}{q_1^2} | \bar{Z}_t \right], \end{aligned}$$

where Q_1 is the normalization factor that is the integral of the target distribution \hat{q}_1 over all the light sub-paths \bar{y}_s with arbitrary length ($s \geq 1$), and $q_1(\bar{y} | \bar{Z}_t)$ is the conditional pdf, and $q_1(\bar{x}) = q_1(\bar{y} | \bar{Z}_t) p(\bar{Z}_t)$. The integral domain \mathcal{A} is a union of A^s as $\mathcal{A} = \cup_{s \geq 1} A^s$. In the similar way, the following relations hold for pdfs $q_{2,s}$ and q_2 , which are defined similarly as $q_{2,s}(\bar{x}) = q_{2,s}(\bar{y}_s | \bar{Z}_t) p(\bar{Z}_t)$ and $q_2(\bar{x}) = q_2(\bar{y} | \bar{Z}_t) p(\bar{Z}_t)$

$$E \left[\frac{f_t^2}{p \hat{q}_2} | \bar{Z}_t \right] E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] = E \left[\frac{f_t^2}{q_{2,s}^2} | \bar{Z}_t \right], \quad \left(\sum_{s \geq 1} E \left[\frac{f_t^2}{p \hat{q}_2} | \bar{Z}_t \right] \right) \left(\sum_{s \geq 1} E \left[\frac{\hat{q}_2}{p} | \bar{Z}_t \right] \right) = E \left[\frac{f_t^2}{q_2^2} | \bar{Z}_t \right],$$

Then the conditional variance $V[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]$ is expressed by:

$$\frac{1}{M_1} \sum_{s \geq 1} V \left[\frac{f_t}{p} | \bar{Z}_t \right] - \frac{1}{M_1 M_2} \sum_{s \geq 1} V \left[\frac{f_t}{q_{1,s}} | \bar{Z}_t \right] + \frac{1}{M_2} V \left[\frac{f_t}{q_1} | \bar{Z}_t \right] - \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \sum_{s \geq 1} V \left[\frac{f_t}{q_{s,2}} | \bar{Z}_t \right] + \frac{1}{N} \left(1 - \frac{1}{M_2} \right) V \left[\frac{f_t}{q_2} | \bar{Z}_t \right]. \quad (11)$$

2.4. Derivation of the variance $V[\langle I_t \rangle_{\text{tsr}}^N]$ for two-stage resampling

We now derive the variance $V[\langle I_t \rangle_{\text{tsr}}^N]$ by taking into account the randomness of the eye sub-path \bar{Z}_t . The variance $V[\langle I_t \rangle_{\text{tsr}}^N]$ is calculated by the law of total variance as:

$$V[\langle I_t \rangle_{\text{tsr}}^N] = E[V[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]] + V[E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]]. \quad (12)$$

So far, we have included the pdf $p(\bar{Z}_t)$ in the pdfs p , $q_{1,s}$, $q_{2,s}$, q_1 , and q_2 . To take into account the randomness of the eye sub-path \bar{Z}_t , we decompose the pdfs into the pdf $p(\bar{Z}_t)$ and the (conditional) pdfs as:

$$p(\bar{x}_s) = p(\bar{y}_s) p(\bar{Z}_t), q_{1,s}(\bar{x}_s) = q_{1,s}(\bar{y}_s | \bar{Z}_t) p(\bar{Z}_t), q_1(\bar{x}) = q_1(\bar{y} | \bar{Z}_t) p(\bar{Z}_t).$$

By using $E[V[a|b]b^2] = V[ab] - V[E[a|b]b]$ and considering $1/p(\bar{Z}_t)$ as random variables, the following expected value of variance is expressed by:

$$E \left[V \left[\frac{f_t(\bar{X}_s)}{p(\bar{X}_s)} | \bar{Z}_t \right] \right] = E \left[V \left[\frac{f_t(\bar{X}_s)}{p(\bar{Y}_s)} | \bar{Z}_t \right] \frac{1}{p^2(\bar{Z}_t)} \right] = V \left[\frac{f_t(\bar{X}_s)}{p(\bar{Z}_t) p(\bar{Y}_s)} \right] - V \left[E \left[\frac{f_t(\bar{X}_s)}{p(\bar{Y}_s)} | \bar{Z}_t \right] \frac{1}{p(\bar{Z}_t)} \right] = V \left[\frac{f_t(\bar{X}_s)}{p(\bar{X}_s)} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{y}_s \bar{Z}_t) d\mu(\bar{y}_s) \right].$$

The expected value of each variance term in Eq. (10) can be rewritten in the similar way. Then $E[V[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]]$ is expressed by:

$$\begin{aligned} & \frac{1}{M_1} \sum_{s \geq 1} \left(V \left[\frac{f_t}{p} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] \right) - \frac{1}{M_1 M_2} \sum_{s \geq 1} \left(V \left[\frac{f_t}{q_{1,s}} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] \right) \\ & + \frac{1}{M_2} \left(V \left[\frac{f_t}{q_1} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{\mathcal{A}} f_t(\bar{x}) d\mu(\bar{y}) \right] \right) - \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \sum_{s \geq 1} \left(V \left[\frac{f_t}{q_{s,2}} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] \right) \\ & + \frac{1}{N} \left(1 - \frac{1}{M_2} \right) \left(V \left[\frac{f_t}{q_2} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{\mathcal{A}} f_t(\bar{x}) d\mu(\bar{y}) \right] \right). \end{aligned}$$

By rearranging the above equation and adding $V[E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]]$, the variance $V[\langle I_t \rangle_{\text{tsr}}^N]$ is represented by:

$$\begin{aligned} & \frac{1}{M_1} \sum_{s \geq 1} \left(V \left[\frac{f_t}{p} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] \right) - \frac{1}{M_1 M_2} \sum_{s \geq 1} \left(V \left[\frac{f_t}{q_{1,s}} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] \right) \\ & + \frac{1}{M_2} \left(V \left[\frac{f_t}{q_1} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{\mathcal{A}} f_t(\bar{x}) d\mu(\bar{y}) \right] \right) - \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \sum_{s \geq 1} \left(V \left[\frac{f_t}{q_{s,2}} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] \right) \\ & + \frac{1}{N} \left(1 - \frac{1}{M_2} \right) \left(V \left[\frac{f_t}{q_2} \right] - V \left[\frac{1}{p(\bar{Z}_t)} \int_{\mathcal{A}} f_t(\bar{x}) d\mu(\bar{y}) \right] \right) + \underbrace{V \left[\frac{1}{p(\bar{Z}_t)} \int_{\mathcal{A}} f_t(\bar{x}) d\mu(\bar{y}) \right]}_{V[E[\langle I_t \rangle_{\text{tsr}}^N | \bar{Z}_t]]} \\ & = \frac{1}{M_1} \sum_{s \geq 1} V \left[\frac{f_t}{p} \right] - \frac{1}{M_1 M_2} \sum_{s \geq 1} V \left[\frac{f_t}{q_{1,s}} \right] + \frac{1}{M_2} V \left[\frac{f_t}{q_1} \right] - \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \sum_{s \geq 1} V \left[\frac{f_t}{q_{s,2}} \right] + \frac{1}{N} \left(1 - \frac{1}{M_2} \right) V \left[\frac{f_t}{q_2} \right] \\ & - \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \left(1 - \frac{1}{N} \right) \sum_{s \geq 1} V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] + \left(1 - \frac{1}{N} \right) \left(1 - \frac{1}{M_2} \right) V \left[\frac{1}{p(\bar{Z}_t)} \int_{\mathcal{A}} f_t(\bar{x}) d\mu(\bar{y}) \right]. \end{aligned}$$

In summary, the variance of the (N -sample) two-stage resampling estimator $V[\langle I_t \rangle_{\text{tsr}}^N]$ for the resampling strategy t is represented by:

$$\begin{aligned} V[\langle I_t \rangle_{\text{tsr}}^N] &= \frac{1}{M_1} \sum_{s \geq 1} V \left[\frac{f_t}{p} \right] - \frac{1}{M_1 M_2} \sum_{s \geq 1} V \left[\frac{f_t}{q_{1,s}} \right] + \frac{1}{M_2} V \left[\frac{f_t}{q_1} \right] - \frac{1}{N} \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \sum_{s \geq 1} V \left[\frac{f_t}{q_{s,2}} \right] + \frac{1}{N} \left(1 - \frac{1}{M_2} \right) V \left[\frac{f_t}{q_2} \right] \\ & - \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \left(1 - \frac{1}{N} \right) \sum_{s \geq 1} V \left[\frac{1}{p(\bar{Z}_t)} \int_{A^s} f_t(\bar{x}_s) d\mu(\bar{y}_s) \right] + \left(1 - \frac{1}{N} \right) \left(1 - \frac{1}{M_2} \right) V \left[\frac{1}{p(\bar{Z}_t)} \int_{\mathcal{A}} f_t(\bar{x}) d\mu(\bar{y}) \right]. \end{aligned} \quad (13)$$

The second line in Eq. (13) includes the variance terms of integrals, which makes the minimization of the upper bound of the variance $V[\langle I_t \rangle_{\text{tsr}}^N]$ infeasible. To address this problem, we set the number of samples to one ($N = 1$) to vanish the coefficients for the variance terms including the integrals as:

$$V[\langle I_t \rangle_{\text{tsr}}^1] = \frac{1}{M_1} \sum_{s \geq 1} V \left[\frac{f_t}{p} \right] - \frac{1}{M_1 M_2} \sum_{s \geq 1} V \left[\frac{f_t}{q_{1,s}} \right] + \frac{1}{M_2} V \left[\frac{f_t}{q_1} \right] - \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \sum_{s \geq 1} V \left[\frac{f_t}{q_{s,2}} \right] + \left(1 - \frac{1}{M_2} \right) V \left[\frac{f_t}{q_2} \right]. \quad (14)$$

3. Derivation of weighting functions w_t for two-stage resampling

3.1. The variance of the pixel measurement $V[\langle I \rangle]$

We now derive the weighting functions by minimizing the upper bound of the variance $V[\langle I \rangle]$ of the pixel measurement I . For a light path \bar{x} with length k , the variance $V[\langle I \rangle]$ is calculated by the sum of variances for $k+2$ strategies as:

$$V[\langle I \rangle] = \sum_{t \in \Lambda_{\text{tsr}}} V[\langle I_t \rangle_{\text{tsr}}^1] + \frac{1}{N_t} \sum_{t \in \Lambda_{\text{is}}} V \left[\frac{f_t}{p} \right],$$

where $\Lambda_{\text{tsr}} \in \{2, \dots, k\}$ represents the resampling strategies and $\Lambda_{\text{is}} \in \{0, 1, k+1\}$ represents the strategies handled by BPT, and N_t is the number of samples for BPT. By substituting $V[\langle I_t \rangle_{\text{tsr}}^1]$ in Eq. (14), $V[\langle I \rangle]$ is expressed by:

$$V[\langle I \rangle] = \sum_{t \in \Lambda_{\text{tsr}}} \left(\frac{1}{M_1} V \left[\frac{f_t}{p} \right] - \frac{1}{M_1 M_2} V \left[\frac{f_t}{q_{1,s}} \right] + \frac{1}{M_2} V \left[\frac{f_t}{q_1} \right] - \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) V \left[\frac{f_t}{q_{s,2}} \right] + \left(1 - \frac{1}{M_2} \right) V \left[\frac{f_t}{q_2} \right] \right) + \frac{1}{N_t} \sum_{t \in \Lambda_{\text{is}}} V \left[\frac{f_t}{p} \right], \quad (15)$$

where the summation over s is eliminated since the number of the light sub-path vertices is uniquely determined as $s = k+1-t$.

3.2. Derivation of the weighting functions

We derive the weighting functions that minimize the upper bound of the variance $V[\langle I \rangle] = E[\langle I^2 \rangle] - E[\langle I \rangle]^2$. Specifically, we aim to minimize the second moment $E[\langle I^2 \rangle]$ as:

$$E[\langle I^2 \rangle] = \sum_{t \in \Lambda_{tsr}} \left(\frac{1}{M_1} E \left[\frac{f_t^2}{p^2} \right] - \frac{1}{M_1 M_2} E \left[\frac{f_t^2}{q_{1,s}^2} \right] + \frac{1}{M_2} E \left[\frac{f_t^2}{q_1^2} \right] - \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) E \left[\frac{f_t^2}{q_{s,2}^2} \right] + \left(1 - \frac{1}{M_2} \right) E \left[\frac{f_t^2}{q_2^2} \right] \right) + \frac{1}{N_t} \sum_{t \in \Lambda_{is}} E \left[\frac{f_t^2}{p^2} \right].$$

We represent each expected value term with the integral form as $E[f^2/p^2] = \int f^2(\bar{x})/p(\bar{x})d\mu(\bar{x})$. Since it is sufficient to minimize the integrand at each path \bar{x} separately and $f(\bar{x})$ is constant for all strategies, we minimize the following objective function subject to the condition $\sum_{t=0}^{k+1} w_t = 1$:

$$\sum_{t \in \Lambda_{tsr}} \left(\frac{1}{M_1} \frac{w_t^2}{p} - \frac{1}{M_1 M_2} \frac{w_t^2}{q_{1,s}} + \frac{1}{M_2} \frac{w_t^2}{q_1} - \frac{1}{M_1} \left(1 - \frac{1}{M_2} \right) \frac{w_t^2}{q_{s,2}} + \left(1 - \frac{1}{M_2} \right) \frac{w_t^2}{q_2} \right) + \frac{1}{N_t} \sum_{t \in \Lambda_{is}} \frac{w_t^2}{p}.$$

In the above equation, we represent the pdfs $q_{1,s}$ and $q_{2,s}$ with q_1 and q_2 as:

$$q_{1,s}(\bar{x}) = \frac{\hat{q}_1(\bar{x})}{\int_{A^s} \hat{q}_1(\bar{x}) d\mu(\bar{y}_s)} = \frac{\hat{q}_1(\bar{x})}{\int_{\mathcal{A}} \hat{q}_1(\bar{x}) d\mu(\bar{y})} \cdot \frac{\int_{\mathcal{A}} \hat{q}_1(\bar{x}) d\mu(\bar{y})}{\int_{A^s} \hat{q}_1(\bar{x}) d\mu(\bar{y}_s)} = q_1(\bar{x}) \frac{Q_1}{Q_{1,s}}, \quad (16)$$

where $Q_{1,s}$ is the normalization factor of $\hat{q}_{1,s}$ over A^s , and Q_1 is that of \hat{q}_1 over \mathcal{A} . Similarly, $q_{2,s}(\bar{x}) = q_2(\bar{x})Q_2/Q_{2,s}$. By using these, the objective function is represented as:

$$\sum_{t \in \Lambda_{tsr}} \left(\frac{1}{M_1} \frac{1}{p(\bar{x})} + \left(1 - \frac{1}{M_1} \frac{Q_{1,s}}{Q_1} \right) \frac{1}{M_2} \frac{1}{q_1(\bar{x})} + \left(1 - \frac{1}{M_1} \frac{Q_{2,s}}{Q_2} \right) \left(1 - \frac{1}{M_2} \right) \frac{1}{q_2(\bar{x})} \right) w_t^2(\bar{x}) + \sum_{t \in \Lambda_{is}} \frac{1}{N_t} \frac{w_t(\bar{x})^2}{p(\bar{x})}.$$

Although it is possible to estimate $Q_{1,s}$ and $Q_{2,s}$ by using Monte Carlo integration, it is difficult to keep the normalization factors $Q_{1,s}$ and $Q_{2,s}$ for all the non-negative integers s . Fortunately, we can bound the ratios $Q_{1,s}/Q_1$ and $Q_{2,s}/Q_2$ between zero and one by definition, and we approximate the ratios $Q_{1,s}/Q_1$ and $Q_{2,s}/Q_2$ with the upper bound one. Then we define the following density p_{tsr} as:

$$p_{tsr}(\bar{x}) = \left(\frac{1}{M_1} \frac{1}{p(\bar{x})} + \left(1 - \frac{1}{M_1} \right) \left(\frac{1}{M_2} \frac{1}{q_1(\bar{x})} + \left(1 - \frac{1}{M_2} \right) \frac{1}{q_2(\bar{x})} \right) \right)^{-1}.$$

By using the density p_{tsr} , the objective function is simplified to $\sum_{t=0}^{k+1} n_t p_t(\bar{x})$ where $n_t = 1$ and $p_t(\bar{x}) = p_{tsr}(\bar{x})$ for $t \in \Lambda_{tsr}$ and $n_t = N_t$ for $t \in \Lambda_{is}$. This makes it possible to derive the weighting function using the balance heuristic as:

$$w_t(\bar{x}) = \frac{n_t p_t(\bar{x})}{\sum_{i=0}^{k+1} n_i p_i(\bar{x})}, \quad (17)$$

$$n_t = \begin{cases} 1 & (t \in \Lambda_{tsr}) \\ N_t & (t \in \Lambda_{is}) \end{cases}, \quad (18)$$

$$p_t(\bar{x}) = \begin{cases} \left(\frac{1}{M_1} \frac{1}{p(\bar{x})} + \left(1 - \frac{1}{M_1} \right) \left(\frac{1}{M_2} \frac{1}{q_1(\bar{x})} + \left(1 - \frac{1}{M_2} \right) \frac{1}{q_2(\bar{x})} \right) \right)^{-1} & (t \in \Lambda_{tsr}) \\ p(\bar{x}) & (t \in \Lambda_{is}). \end{cases} \quad (19)$$